

CLEARINGHOUSE FOR FEDERAL SCIENTIFIC AND TECHNICAL INFORMATION CFSTI
DOCUMENT MANAGEMENT BRANCH 410.11

LIMITATIONS IN REPRODUCTION QUALITY

ACCESSION #

604222

- ☒ 1. WE REGRET THAT LEGIBILITY OF THIS DOCUMENT IS IN PART UNSATISFACTORY. REPRODUCTION HAS BEEN MADE FROM BEST AVAILABLE COPY.
- ☐ 2. A PORTION OF THE ORIGINAL DOCUMENT CONTAINS FINE DETAIL WHICH MAY MAKE READING OF PHOTOCOPY DIFFICULT.
- ☐ 3. THE ORIGINAL DOCUMENT CONTAINS COLOR, BUT DISTRIBUTION COPIES ARE AVAILABLE IN BLACK-AND-WHITE REPRODUCTION ONLY.
- ☐ 4. THE INITIAL DISTRIBUTION COPIES CONTAIN COLOR WHICH WILL BE SHOWN IN BLACK-AND-WHITE WHEN IT IS NECESSARY TO REPRINT.
- ☐ 5. LIMITED SUPPLY ON HAND: WHEN EXHAUSTED, DOCUMENT WILL BE AVAILABLE IN MICROFICHE ONLY.
- ☐ 6. LIMITED SUPPLY ON HAND: WHEN EXHAUSTED DOCUMENT WILL NOT BE AVAILABLE.
- ☐ 7. DOCUMENT IS AVAILABLE IN MICROFICHE ONLY.
- ☐ 8. DOCUMENT AVAILABLE ON LOAN FROM CFSTI (TT DOCUMENTS ONLY).
- ☐ 9.

PROCESSOR: *PM*

604222

604222

1

A SIMPLE DERIVATION OF
THE POISSON DISTRIBUTION

R. E. Kalaba

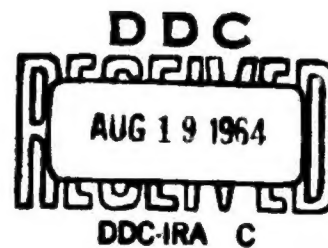
P-414 ✓

11 June 1953

Approved for OTS release

3 p

COPY	1	OF	1
HARD COPY	\$.1.00		
MICROFICHE	\$.0.50 p		



A Simple Derivation of the Poisson Distribution

One of the most important stochastic processes is the Poisson process, in which it is assumed that (a) the numbers of events occurring in nonoverlapping time intervals are independent; (b) the probability of one event's occurring during time dt is $\lambda dt + o(dt)$, where λ is a constant, while the probability that two or more occur is $o(dt)$. Various approaches [1] are known which lead to the result that the probability that n events occur in time t , $p_n(t)$, is

$$(1) \quad p_n(t) = e^{-\lambda t} \frac{(\lambda t)^n}{n!} \quad (n = 0, 1, 2, \dots),$$

that is, the Poisson distribution with mean λt . An extremely simple and straightforward derivation of this formula, based on an idea of G. Morant [2], is as follows:

The probability for no event to occur in time t is

$$(2) \quad p_0(t) = \lim_{dt \rightarrow 0} [1 - \lambda dt - o(dt)]^{t/dt} \\ = e^{-\lambda t}.$$

Eq. (2) merely expresses the fact that for no event to occur in time t , none may occur in any of the subintervals of length dt into which t may be divided. We now use this result to aid in obtaining Eq. (1).

Let us consider n small nonoverlapping time intervals dt_1, dt_2, \dots, dt_n , contained within the time interval $(0, t)$. The probability that n events occur—the first at time t_1 within the interval dt_1 , the second at time t_2 within the interval dt_2 , and so on—is asymptotically equal to

$$e^{-\lambda t_1} \lambda dt_1 e^{-\lambda(t_2-t_1)} \lambda dt_2 \dots \lambda dt_n e^{-\lambda(t-t_n)},$$

which reduces to

$$e^{-\lambda t} \lambda^n dt_1 dt_2 \dots dt_n.$$

This obtains since no event occurs from time 0 to time t_1 ; one event occurs in the interval dt_1 ; no event occurs from time t_1 to t_2 ; and so on. Hence $p_n(t)$, which is the integral of this expression over all t_n satisfying

$$0 \leq t_1 \leq t_2 \leq \dots \leq t_n \leq t,$$

is given by

$$(3) \quad p_n(t) = e^{-\lambda t} \lambda^n \int_0^t \int_0^{t_n} \dots \int_0^{t_2} dt_1 dt_2 \dots dt_n \quad (n = 1, 2, \dots),$$

which immediately yields Eq. (1), since the integral in Eq. (3) equals $t^n/n!$.

Thus, using only the simplest kind of reasoning from probability theory, we have deduced the Poisson distribution from the basic assumptions (a) and (b). Consequently, the need for viewing the Poisson distribution as a limiting case of some other distribution is obviated. In addition the technique used readily generalizes to the case in which λ depends on t .

References

- [1] W. Feller, An Introduction to Probability Theory and Its Applications, John Wiley and Sons, New York, 1950.
- [2] G. Morant, "On Random Occurrences in Space and Time, When Followed by a Closed Interval," Biometrika, Vol XIII (1921), pp 309-337.